Estimation of parameters of a partial neutrino mixing matrix

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## Chapter 1

## Introduction

The partial neutrino mixing matrix predicts the three neutrino mixing angles in terms of two parameters. These two parameters can be constrained using the experimental value of the three neutrino mixing angles. A chi-square analysis is used to extract the value of these two parameters. Moreover, a Monte Carlo study is used to check accuracy of the result.

### 1.1 Mixing Matrix

In 1968 during an experiment at fome-stake it fiad been observed that the flux of electron neutrinos coming from sun was only about one third of that predicted by Behcall. The so called 'Solar neutrino problem'. Solution of this problem was purposed by $\mathcal{B}$. Pontecorvo that electron neutrinos produced by sun are transformed in flight into different species, to which fome-stake experiment was insensitive. Later from experiments like Super-kamiokande and $\mathcal{S N} \mathcal{N O}$ it was confirmed that electron neutrinos
were transforming into muon neutrinos. This is the mechanism we now call neutrino oscillation.

In atmosphere neutrinos are produced from decay of pions and muons.

$$
\begin{array}{ll}
\pi^{+} \rightarrow \mu^{+}+\nu_{\mu}, & \mu^{+} \rightarrow e^{+}+\nu_{e}+\overline{\nu_{\mu}} \\
\pi^{-} \rightarrow \mu^{-}+\overline{\nu_{\mu}}, & \mu^{-} \rightarrow e^{-}+\overline{\nu_{e}}+\nu_{\mu}
\end{array}
$$

Both these reactions produce muon neutrino whereas electron neutrino is produced only in one reaction. Evidently it was supposed that there should be twice as many muon neutrino as electron neutrino. Indeed this was the case for neutrinos coming directly from overfiead which travell a distance of $10 \mathcal{K}$ in or so. As the zenith angle increases and with this distance from the source, ratio of muon to electron neutrino decreases. This suggest muon neutrinos are also changing into tau neutrino.
$\mathcal{N}$ (eutrinos interact as flavour eigenstate $\left(\nu_{e}, \nu_{\mu}, \nu_{\tau}\right)$ and they travel as eigenstates of free particle $\mathcal{H a m i l t o n i a n}$ which are $\left(\nu_{e}, \nu_{\mu}, \nu_{\tau}\right)$ called mass eigenstates. The theory is similar to quantum mechanics of mixed states. flavour eigenstates are linear combination of mass eigenstates.

$$
\begin{align*}
\nu_{e} & =U_{11} \nu_{1}+U_{12} \nu_{2}+U_{13} \nu_{3} \\
\nu_{\mu} & =U_{21} \nu_{1}+U_{22} \nu_{2}+U_{23} \nu_{3}  \tag{1.1}\\
\nu_{\tau} & =U_{31} \nu_{1}+U_{32} \nu_{2}+U_{33} \nu_{3} \tag{1.2}
\end{align*}
$$

These equations can also be written in matrix form:-

$$
\left(\begin{array}{c}
\nu_{e} \\
\nu_{\mu} \\
\nu_{\tau}
\end{array}\right)=\left(\begin{array}{ccc}
U_{11} & U_{12} & U_{13} \\
U_{21} & U_{22} & U_{23} \\
U_{31} & U_{32} & U_{33}
\end{array}\right)\left(\begin{array}{c}
\nu_{1} \\
\nu_{2} \\
\nu_{3}
\end{array}\right)
$$

Matrix on right fand side is called Mixing matrix. General way of constructing this matrix is: -

$$
U=R_{23}\left(\theta_{23}\right) R_{13}\left(\theta_{13}, \delta\right) R_{12}\left(\theta_{12}\right)
$$

Where,

$$
\begin{gathered}
R_{23}=\left(\begin{array}{ccc}
1 & 0 & 0 \\
0 & c_{23} & s_{23} \\
0 & -s_{23} & c_{23}
\end{array}\right), \quad R_{13}=\left(\begin{array}{ccc}
c_{13} & 0 & s_{13} * e^{-\iota \delta} \\
0 & 1 & 0 \\
-s_{13} * e^{\iota \delta} & 0 & c_{13}
\end{array}\right) \\
R_{12}=\left(\begin{array}{ccc}
c 12 & s 12 & 0 \\
s 12 & c 12 & 0 \\
0 & 0 & 1
\end{array}\right)
\end{gathered}
$$

such that

$$
U=\left(\begin{array}{ccc}
c_{12} c_{13} & s_{12} c_{13} & s_{13} e^{-\iota \delta} \\
-s_{12} c_{23}-c_{12} s_{23} s_{13} e^{\iota \delta} & c_{12} c_{23}-s_{12} s_{23} s_{13} e^{\iota \delta} & s_{23} c_{13} \\
s_{12} s_{23}-c_{12} c_{23} s_{13} e^{\iota \delta} & -c_{12} s_{23}-s_{12} c_{23} s_{13} e^{\iota \delta} & c_{23} c_{13}
\end{array}\right)
$$

In general case a mixing matrix fas four parameters, three mixing angles and a $C P$ violating phase. However we can reduce free parameters by using flavour symmetries.
$\mathcal{F l a v o u r ~ s y m m e t r i e s ~ r e d u c e ~ f r e e ~ p a r a m e t e r s ~ e i t h e r ~ b y ~ r e l a t i n g ~ o r ~ p r e d i c t i n g ~ s o m e ~ o f ~}$ them. We are studying the partial mixing matrix whose one column is determine by symmetries and other two colkumns are fixed by unitary conditions. e.g

$$
U=\left(\begin{array}{ccc}
a & * & * \\
b & * & * \\
c & * & *
\end{array}\right) \quad \text { Where } a=2, b=c=1
$$

Our $\mathcal{U}$ is a function of two parameters $(\theta$ and $\phi)$

### 1.2 Statistical Methods

### 1.2.1 Error Propagation

We often want to determine a dependent variable q that is a function of one or more different measured variables. Experientially, variables are often measured with uncertainties. Then how these uncertainties propagate into the dependent variable q?

Case I

Suppose we measured a quantity $x$ and proceed to calculate the quantity

$$
q=x+A
$$

Where $\mathcal{A}$ is a fixed number with no uncertainty. Suppose also that the measurements of $x$ are normally distributed about true valuie $X$ with width $\sigma x$

Probability of obtaining any value $x$ in the interval dx is :

$$
e^{\frac{-(x-X)^{2}}{2 \sigma_{x}^{2}}}
$$

Our problem is to obtain the probability of any value $q=\chi+\mathcal{A}$, or $\chi=q-\mathcal{A}$
probability of obtaining value $\chi=$ Probability of obtaining value $\chi=q-\mathcal{A}$

$$
e^{\frac{-[q-(X+A)]^{2}}{2 \sigma_{x}^{2}}}
$$

This result shows that calculated values of $q$ are centred at $\mathcal{X}+\mathcal{A}$ with uncertainty $\sigma_{x}$

Case II

This time consider quantity to be calculated as

$$
q=A x \quad \text { whiere } A \text { is a fixed number. }
$$

Proceeding same as before we can say that
Probability of obtaining value $q=$ Probability of obtaining value $\chi=q / 6$

$$
\begin{equation*}
e^{\frac{-(q-B X)^{2}}{2 B^{2} \sigma_{x}^{2}}} \tag{1.3}
\end{equation*}
$$

In other words the values of $q=\mathcal{B X}$ will be normally distributed, with center at $q=\mathcal{B} X$ and uncertainty $B \sigma_{x}$

General Case

Suppose we want to measure two independent quantities $x$ and $y$ whose observed values are normally distributed, and we now calculate quantity $q(x, y)$. We
are making an assumption that there uncertainties $\sigma_{x}$ and $\sigma_{y}$ are very small as compared to central value $X$ and $\mathscr{Y}$. This assumption means that we are concerned with only that values of $x$ and $y$ which are very close to $X$ and $\mathscr{S}$. Then we can write $q(x, y)$ as

$$
q(x, y) \approx q(X, Y)+\left(\frac{\partial q}{\partial x}\right) \delta x+\left(\frac{\partial q}{\partial y}\right) \delta y
$$

This approximation is good because most of the time values occur close to central values $X$ and $\mathcal{Y}$. Two partial derivatives are fixed numbers because they are evaluated at $X$ and $\mathcal{Y} \quad \mathcal{H e r e}$ first term is simply a constant term so it mainly shifts the distribution of answers. The second number is the fixed number $\frac{\partial q}{\partial x}$ times $\delta x$ whose distribution fias width (or uncertainty in $\chi$ ) $\sigma x$, so the values of second term are centred on zero with width

$$
\left(\frac{\partial q}{\partial x}\right) \delta x
$$

similarly the values of third term are centred on zero with width

$$
\left(\frac{\partial q}{\partial y}\right) \delta y
$$

combining above three equations we can say that values of $q(x, y)$ are normally distributed about the true value $q(X, Y)$ with width

$$
\sigma_{q}=\sqrt{\left(\frac{\partial q}{\partial x} \delta x\right)^{2}+\left(\frac{\partial q}{\partial y} \delta y\right)^{2}}
$$

If we identify the standard deviation $\sigma_{x}$ and $\sigma_{y}$ as uncertainties in $x$ and $y$ then above result is precisely the rule for propagation of random errors for the case when $q$
is function of just two variables. If $q$ depends on several variables then this equation can be generalised easily as

$$
\sigma_{q}=\sqrt{\left(\frac{\partial q}{\partial x} \delta x\right)^{2}+\left(\frac{\partial q}{\partial y} \delta y\right)^{2} \ldots \ldots}
$$

### 1.2.2 $\chi^{2}$

Suppose we perform some experiment for which we know the distribution of experimental results. We perform experiment several times to record our observations, now question is fow can we decide whether our observed values are in agreement with theoretical values. This can be done by $\chi^{2}$ test. Consider during experiment we get values of dependent variables $q_{i}$. Theoretically $q_{i}$ 's are functions of two unknown variables $x$ and $y$. If $O_{i}$ and $E_{i}$ represents observed and expected values respectively then $\chi^{2}$ can be given as

$$
\chi^{2}=\sum_{i=1}^{n}\left(\frac{p_{i}-f_{p_{i}}(x, y)}{d p_{i}^{\exp }}\right)^{2}
$$

In numerator Difference signifies that we are comparing observed and expected value of $p_{i}$. Square is used so that we always get positive value of difference because if we do not take square then difference can be negative and negative errors can cancel other errors leading to underestimate of errors. It can happen that during experiment that we measure $p_{1}$ with very large accuracy but some other variable $p_{2}$ with small accuracy. Then in $\chi^{2}$ weightage of term including $p_{1}$ should be greater then that of $p_{2}$. That is why we divide the numerator by $d p_{i}$
$\mathcal{D u r i n g}$ estimation of parameters (Here $\chi$ and $y$ ) $\mathcal{W e}$ replace $E_{i}$ with expression for $E_{i}$ then we plot $\chi^{2}$ either contour plot or $3 \mathcal{D}$ plot to know minimum value of $\chi^{2}$.

Minimum value of $\chi^{2}$ curve corresponds to best fit values of $x$ and $y$.

## Monte Carlo

Practically $\mathcal{M o n t e}$ Carlo provides a mathod of simulating experiment and creating models of experimental data. With a Monte Carlo calculation, we can test the statistical significance of data with relative simple calculations that require neither a deep theoretical understanding of statistical analysis nor sopfisticated programming technique.

We have three equations of mixing angles in terms of two known variables $\theta$ and $\phi$. We know there experimental values and uncertainties in them. We also know the distribution followed by $\theta$ and $\phi$. We will generate 10000 values of $\theta$ and $\phi$ randomly around there central value. These values when inserted into expressions of mixing angles we will get random values of mixing angles. $\mathcal{B y}$ making fistogram of these values of mixing angles we can know standard deviation and central values of mixing angles.

## Chapter 2

## Results

2.0.3 Estimation of $\theta$ and $\phi$ using $\chi^{2}$

$$
\begin{array}{ll}
\sin ^{2} \theta_{13}= & \frac{\sin ^{2} \theta}{3} \\
\sin ^{2} \theta_{12}= & 1-\frac{a^{2}}{\left(a^{2}+b^{2}+c^{2}\right) \cos \theta_{13}} \\
\sin ^{2} \theta_{23}= & \frac{1}{2}+\frac{\left(a b c \sin 2 \theta \sqrt{a^{2}+b^{2}+c^{2}}\right)}{\left(b^{2}+c^{2}\right)\left(a^{2}+\left(b^{2}+c^{2}\right) \cos ^{2} \theta\right)} \cos \phi \tag{2.3}
\end{array}
$$

Experimental Values of mixing angles are

- $\sin ^{2} \theta_{13}=0.026 \pm 0.003$
- $\sin ^{2} \theta_{12}=0.32 \pm 0.01$
- $\sin ^{2} \theta_{23}=0.49 \pm 0.08$
and $\chi^{2}$ expression is:-

$$
\chi^{2}=\left(\frac{0.026-\sin ^{2} \theta_{13}}{0.003}\right)^{2}+\left(\frac{0.49-\sin ^{2} \theta_{23}}{0.08}\right)^{2}+\left(\frac{0.32-\sin ^{2} \theta_{12}}{0.01}\right)^{2}
$$



Figure 2.1: Here black region represent the minimized value of $\theta$ and $\phi$. Region outside this is contour corresponding to $1 \sigma$. Minimum numerical value of $\chi^{2}$ is 0.07 corresponding to $\theta=0.28 \pm 0.03$ and $\phi=1.61 \pm 0.6$. Width and height of $1 \sigma$ contour gives us uncertainties for central values of $\theta$ and $\phi$
$a=\sqrt{\frac{2}{3}} ;$
$b=\frac{1}{\sqrt{6}} ;$
$c=\frac{1}{\sqrt{6}} ;$
$\operatorname{ts13}\left[\theta_{-}\right]:=\frac{\operatorname{Sin}[\theta]^{2}}{3}$
$\operatorname{chis} 13\left[\theta_{-}\right]:=\left(\frac{0.026-\operatorname{ts} 13[\theta]}{0.004}\right)^{2}\left(* \operatorname{Chi}\right.$ Square Value of $\left.\sin ^{2} \theta_{13} *\right)$
(*Calculation of Chi Square for $\operatorname{Sin}^{2} \theta_{23} *$ )
$t A\left[\theta_{-}\right]:=\frac{a * b * c * \operatorname{Sin}[2 \theta] \sqrt{a^{2}+b^{2}+c^{2}}}{\left(b^{2}+c^{2}\right)\left(a^{2}+\left(b^{2}+c^{2}\right) \operatorname{Cos}[\theta]^{2}\right)}$
$t s 23\left[\theta_{-}, \phi_{-}\right]:=0.5+\mathrm{tA}[\theta] * \operatorname{Cos}[\phi]$
$\operatorname{chis} 23\left[\theta_{-}, \phi_{-}\right]:=\left(\frac{\operatorname{ts} 23[\theta, \phi]-0.49}{0.08}\right)^{2}\left(* \operatorname{Chi}\right.$ Square for $\left.\sin ^{2} \theta_{23} *\right)$
(*Calculation of Chi Square for $\operatorname{Sin}^{2} \theta_{12} *$ )
$\operatorname{t\theta 13}\left[\theta \_\right]:=\operatorname{ArcSin}[\sqrt{\operatorname{ts} 13[\theta]}]$
$\operatorname{t\theta 12}\left[\theta \_\right]:=\operatorname{ArcCos}\left[\frac{a * \operatorname{Sec}[t \theta 13[\theta]]}{\sqrt{a^{2}+b^{2}+c^{2}}}\right]$
$\operatorname{tsin} 12\left[\theta_{-}\right]:=\operatorname{Sin}[\operatorname{tel} 2[\theta]]^{2}$
$\operatorname{chis12}\left[\theta_{-}\right]:=\left(\frac{0.32-t \sin 12[\theta]}{0.017}\right)^{2}$
(*Final Chi Squared*)
chisquare $\left[\theta_{-}, \phi_{-}\right]:=\operatorname{chis} 13[\theta]+\operatorname{chis} 23[\theta, \phi]+\operatorname{chis} 12[\theta]$
Plot3D [chisquare $[\theta, \phi],\{\theta, 0.17,0.35\},\{\phi,-3,3\}]$
kk1 $=$ ContourPlot [chisquare $[\theta, \phi],\{\theta, 0.17,0.35\}$,
$\{\phi,-3,3\}$, Contours $\rightarrow\{2.30,4.61,9.21\}$, FrameLabel $\rightarrow$ Automatic];
FindMinimum [chisquare $[\theta, \phi],\{\{\theta, 0.2\},\{\phi, 1.5\}\}]$ (*Minimum Value of $\theta$ and $\phi *$ ) kk2 $=$ ContourPlot [chisquare $[\theta, \phi]=0.0779592$,
$\{\theta, 0.17,0.35\},\{\phi,-3,3\}$, ContourStyle $\rightarrow$ Directive[Black, Thick]];
Show [
kk1,
kk2]

### 2.1 Error Propagation

We can calculate errors in mixing angles by using expressions previously discussed in 1.2.1 to get

$$
\begin{gather*}
\sigma_{13}=\frac{\sin 2 \theta}{3} \delta \theta \\
\sigma_{12}=  \tag{2.4}\\
\text { where } \delta \theta_{13}=  \tag{2.5}\\
\text { and } N=  \tag{2.6}\\
\sigma_{23}=  \tag{2.7}\\
\text { whiere } \delta A=\left(\frac{\cos \theta}{\sqrt{3} \sqrt{1-\frac{\sin ^{2} \theta}{3}}} \delta \theta\right.  \tag{2.8}\\
\frac{\sec ^{2} \theta_{13} \tan \theta_{13} \delta \theta_{13}}{\sqrt{a^{2}+b^{2}+c^{2}}}  \tag{2.9}\\
\text { and } \delta \cos \phi= \\
\left.\frac{\sqrt{\frac{2}{3}} \cos [2 \theta]}{\frac{2}{3}+\frac{\cos [\theta]^{2}}{3}}+\frac{\sqrt{\frac{2}{3}} \cos [\theta] \sin [\theta] \sin [2 \theta]}{3\left(\frac{2}{3}+\frac{\cos [\theta]^{2}}{3}\right)^{2}}\right) \delta \theta \\
\sin \phi \delta \phi
\end{gather*}
$$

Experimental values of mixing angles:-

- $\sin ^{2} \theta_{13}=0.026 \pm 0.004$
- $\sin ^{2} \theta_{12}=0.32 \pm 0.01$
- $\sin ^{2} \theta_{23}=0.49 \pm 0.08$

Values of mixing angles from error propagation: -

- $\sin ^{2} \theta_{13}=0.026 \pm 0.005$
- $\sin ^{2} \theta_{12}=0.32 \pm 0.01$
- $\sin ^{2} \theta_{23}=0.49 \pm 16.37$

Here error in $\sin ^{2} \theta_{23}$ is very large because in expression of error propagation for $\sin ^{2} \theta_{23}$ there is $\cos \phi$ term in denominator which goes to inf as $\phi$ goes to 90 degree.
(*Error Propagation in Mathematica 8*)
Quit []
$a=\sqrt{\frac{2}{3}} ;$
$b=\frac{1}{\sqrt{6}}$;
$c=\frac{1}{\sqrt{6}}$;
$\theta=0.2838 ;$
$\delta \theta=0.03 ;$
$\phi=-1.61$;
$\delta \phi=1.3$;
Print["True Value of $\operatorname{Sin} \theta^{2}{ }_{13}$ is : - "]
$\operatorname{ts} 13=\frac{\operatorname{Sin}[\theta]^{2}}{3}\left(*\right.$ True Value of $\left.\operatorname{Sin} \theta^{2}{ }_{13} *\right)$
Print["Error in the value of $\operatorname{Sin} \theta^{2}{ }_{13}$ is : - "]
es13 $=\operatorname{Sin}[2 \theta] * \delta \theta$
(*fes $13=\frac{2 * \operatorname{Cos}[\theta]}{3 * \operatorname{Sin}[\theta]} * \delta \theta$; (*Error in $\left.\operatorname{Sin}^{2} \theta_{13} *\right)$
es13= fes13*ts13;*)
$t A=\frac{a * b * c * \sin [2 \theta] \sqrt{a^{2}+b^{2}+c^{2}}}{\left(b^{2}+c^{2}\right)\left(a^{2}+\left(b^{2}+c^{2}\right) \operatorname{Cos}[\theta]^{2}\right)} ;(*$ True Value of $A *$
Print ["True Value of $\operatorname{Sin} \theta^{2}{ }_{23}$ is : - "]
$t s 23=\frac{c^{2}}{b^{2}+c^{2}}+t A * \operatorname{Cos}[\phi]\left(*\right.$ True Value of $\left.\operatorname{Sin} \theta_{23} *\right)$
$e c \phi=\operatorname{Sin}[\phi] * \delta \phi ;$
numeratort $A=a * b * c * \operatorname{Sin}[2 \theta] \sqrt{a^{2}+b^{2}+c^{2}}$;
denominatort $A=\left(b^{2}+c^{2}\right)\left(a^{2}+\left(b^{2}+c^{2}\right) \operatorname{Cos}[\theta]^{2}\right)$;
numeratore $A=a * b * c * 2 * 2 * \operatorname{Cos}[2 \theta] * \sqrt{a^{2}+b^{2}+c^{2}} * \delta \theta$;
denominatore $=2 * \operatorname{Sin}[\theta] * \delta \theta\left(b^{2}+c^{2}\right)^{2}$;
frace $\mathbf{A}=\frac{\text { numeratore } \mathbf{A}}{\text { numeratort } \mathbf{A}}+\frac{\text { denominatore } \mathbf{A}}{\text { denominatort } \mathbf{A}}$;
eA = fraceA * tA;

### 2.2 Monte Carlo



Here in all the fistograms we get same best fit and uncertainties of mixing angles as given by experiments which verifies our result of $\chi^{2}$

```
\(a=\sqrt{\frac{2}{3}} ;\)
b \(=\frac{1}{\sqrt{6}}\);
\(c=\frac{1}{\sqrt{6}}\);
ttheta := RandomVariate [NormalDistribution [16.5 \(\circ\), \(1{ }^{\circ}\) ]]
pphi := RandomVariate [NormalDistribution [-91.85 \(\circ\), \(20^{\circ}\) ]]
ss13 : \(=\frac{\operatorname{Sin}[\text { ttheta }]^{2}}{3}\);
(* \(\left.\operatorname{Sin}_{13}{ }^{2} *\right)\)
ttheta13 \(:=\operatorname{ArcSin}[\sqrt{\mathbf{s s 1 3}}] ;\left(* \theta_{13} *\right)\)
    \(a^{2}\)
\(\boldsymbol{s s} 12:=1-\overline{\left(a^{2}+b^{2}+c^{2}\right) \operatorname{Cos}[t \text { theta13] }} ;\left(* \operatorname{Sin}_{12}{ }^{2} *\right)\)
\(\boldsymbol{s s} 23:=0.5+\frac{\left(a * b * c * \sin [2 * \text { theta }] \sqrt{a^{2}+b^{2}+c^{2}}\right)}{\left(b^{2}+c^{2}\right)\left(a^{2}+\left(b^{2}+c^{2}\right) \operatorname{Cos}[\text { ttheta }]^{2}\right)} * \operatorname{Cos}[p p h i]\)
Histogram [Labeled [Table [ss13, \{j, 1, 1000\}], \(\operatorname{Sin}^{2} \theta_{13} "\), Above],
ChartElementFunction \(\rightarrow\) "FadingRectangle", ChartStyle \(\rightarrow\) Orange]
Histogram [Labeled [Table [ss12, \{j, 1, 1000\}], " \(\operatorname{Sin}^{2} \theta_{12} "\), Above],
ChartElementFunction \(\rightarrow\) "FadingRectangle", ChartStyle \(\rightarrow\) Red]
Histogram [Labeled [Table [ss23, \{j, 1, 1000\}], \(\operatorname{Sin}^{2} \theta_{23} "\), Above],
ChartElementFunction \(\rightarrow\) "FadingRectangle", ChartStyle \(\rightarrow\) Blue]
n = 10000 ;
Print ["Mean Value of \(\operatorname{Sin}^{2} \theta_{12}\) "]
mss12 = Mean [Table[ss12, \{j, 1, n\}] ]
Print ["Mean Value of \(\operatorname{Sin}^{2} \theta_{23}\) "]
mss23 \(=\) Mean [Table [ss23, \{j, 1, n\}]]
Print ["Mean value of \(\operatorname{Sin}^{2} \theta_{13}\) "]
mss13 = Mean [Table [ss13, \{j, 1, n\}]]
Print ["Standard Deviation of \(\operatorname{Sin}^{2} \theta_{12}\) "]
StandardDeviation [Table[ss12, \{j, 1, n\}]]
Print ["Standard Deviation of \(\operatorname{Sin}^{2} \theta_{23}\) "]
StandardDeviation [Table[ss23, \{j, 1, n\}]]
```


## Chapter 3

## Summary

We studied how to estimate parameters using $\chi^{2}$ which we then used to estimate two variables from experimental values of three mixing angles. These results are verified first using error propagation and then Monte Carlo method. Uncertainties for $\sin ^{2} \theta_{12}$ given by error propagation are not matching with experimental values because error propagation assumes that uncertainty should be very small as compared to central value but in case of $\sin ^{2} \theta_{12}$ uncertainty becomes infinite as $\phi$ approaches 90 degree. $\mathcal{H o w e v e r ~ r e s u l t s ~ p r o v i d e d ~ b y ~ M o n t e ~ C a r l o ~ a r e ~ i n ~ e x a c t ~ m a t c h ~ w i t h ~ e x p e r i m e n t a l ~ v a l u e s ~}$ verifying our result of $\chi^{2}$.

